

# A MODIFIED RATIO ESTIMATOR BASED ON THE COEFFICIENT OF VARIATION IN DOUBLE SAMPLING

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## SUMMARY

Utilising the value of coefficient of variation for an auxiliary variable, a modified ratio estimator for double sampling is proposed which is more efficient than the usual ratio estimator in double sampling and the simple mean estimator when  $\rho_{yx}$  lies between certain range.

## INTRODUCTION

Sisodia and Dwivedi [2] noted that by utilising the value of  $C_x$ , coefficient of variation (C.V.) for an auxiliary variable, a modified ratio estimator is more efficient than both ratio and simple mean estimator when  $\rho_{yx}$  lies within a certain range. When information on population mean  $\bar{X}_N$  is not available double sampling is used. The usual biased ratio estimator in double sampling is given by

$$\hat{Y}_{RD} = \frac{\bar{y}_n}{\bar{x}_n} \bar{x}_n, \quad \dots(I.1)$$

and its bias and the mean square error are given by Sukhatme and Sukhatme [3].

In the present paper utilising the value of  $C_x$ , a modified ratio estimator in double sampling is proposed and its efficiency has been compared with the corresponding ratio estimator and the simple mean estimator. A cost function is also considered and the optimum values of  $n'$  and  $n$ , and the optimum variance of the proposed estimator are obtained. The efficiency of the proposed estimator has been illustrated with the example given by Jessen [1].

## MODIFIED RATIO ESTIMATOR IN DOUBLE SAMPLING

The proposed estimator is

$$\hat{Y}_{MRD} = \frac{\bar{y}_n(\bar{x}'_n + C_x)}{(\bar{x}_n + C_x)} \quad \dots(2.1)$$

It can be easily seen that to the first order of approximation

$$E(\hat{Y}_{MRD}) = \bar{Y}_N \left\{ 1 + \left( \frac{1}{n} - \frac{1}{n'} \right) (C_x'^2 - \rho_{yx} C_y C_x') \right\} \quad \dots(2.2)$$

Where

$$C_x' = S_x/\bar{X}'_N, \rho_{yx} = S_{yx}/S_x S_y, C_y = S_y/\bar{Y}_N, \bar{X}'_N = \bar{X}_N + C_x$$

With its relative bias given by

$$B^* = \left( \frac{1}{n} - \frac{1}{n'} \right) (C_x'^2 - \rho_{yx} C_y C_x') \quad \dots(2.3)$$

and MSE given by

$$M^* = \left( \frac{1}{n} - \frac{1}{n'} \right) \left\{ S_y^2 + \frac{\bar{Y}_N^2}{\bar{X}'_N{}^2} S_x^2 - 2 \frac{\bar{Y}_N}{\bar{X}'_N} S_{yx} \right\} + \left( \frac{1}{n'} - \frac{1}{N} \right) S_y^2 \dots(2.4)$$

If  $n' = N$ , the  $MSE$  of  $\hat{Y}_{MRD}$  to the first order of approximation reduces to  $MSE$  of modified ratio estimator proposed by Sisodia and Dwivedi (1981). When the information about  $C_x$  is not used we get  $\bar{X}'_N = \bar{X}_N$ , the  $MSE$  of proposed estimator reduces to the  $MSE$  of the usual ratio estimator in double sampling.

COMPARISON OF  $\hat{Y}_{MRD}$  WITH THE USUAL RATIO ESTIMATOR  
IN THE DOUBLE SAMPLING AND WITH THE SIMPLE MEAN  
ESTIMATOR :

The estimator  $\hat{Y}_{MRD}$  will be more efficient than  $\hat{Y}_{RD}$  if,  
 $M > M^*$

$$i.e. R_N^2 S_X^2 - 2R_N \rho_{yx} S_x S_y > \left( \frac{\bar{Y}_N}{\bar{X}'_N} \right)^2 S_x^2 - 2 \left( \frac{\bar{Y}_N}{\bar{X}'_N} \right) S_x S_y \rho_{yx}$$

$$i.e. \rho_{yx} < \frac{1}{2} \frac{C_x}{C_y} \left( \frac{2\bar{X}_N + C_x}{\bar{X}_N + C_x} \right) \quad \dots(3.1)$$

The estimator  $\hat{Y}_{MRD}$  will be more efficient than the simple mean estimator if

$$V(\hat{Y}_n) > M^*$$

$$i.e. \left(\frac{1}{n} - \frac{1}{N}\right) S_y^2 > \left(\frac{1}{n} - \frac{1}{n'}\right) \left\{ S_y^2 + \left(\frac{Y_N}{\bar{X}_N}\right)^2 S_x^2 - 2 \left(\frac{Y_N}{\bar{X}_N}\right) \rho_{yx} S_x S_y \right\} + \left(\frac{1}{n'} - \frac{1}{N}\right) S_y^2$$

$$i.e. \rho_{yx} > \frac{1}{2} C'_x / C_y \quad \dots(3.2)$$

Combining (3.1) and (3.2) we find that the modified ratio estimator in double sampling is more efficient than the ratio estimator in double sampling and simple mean estimator if

$$\frac{1}{2} \frac{C_x}{C_y \left(1 + \frac{C_x}{\bar{X}_N}\right)} < \rho_{yx} < \frac{1}{2} \frac{C_x (2\bar{X}_N + C_x)}{C_y (\bar{X}_N + C_x)}$$

**COST FUNCTION**

Consider the simple cost function

$$C = n' C' + n C'' \quad \dots(4.1)$$

Where  $C'$  and  $C''$  are the unit costs of observing an  $X_i$  [in phase I] and  $Y_i$  [in phase II] respectively and  $C$  is the total budget. It can be shown that

$$Opt. Var. (\hat{Y}_{MRD}) = \frac{(\sqrt{C'' S_1'^2} + \sqrt{C'' S_1''^2})^2}{C} \quad \dots(4.2)$$

where

$$S_1'^2 = 2 \left(\frac{Y_N}{\bar{X}_N}\right) S_{yx} - \left(\frac{Y_N}{\bar{X}_N}\right)^2 S_x^2$$

$$S_1''^2 = S_y^2 - 2 \left(\frac{Y_N}{\bar{X}_N}\right) S_{yx} + \left(\frac{Y_N}{\bar{X}_N}\right)^2 S_x^2$$

$N$	$C$	$C''$	$C'$	$\bar{y}_n$	$\bar{x}_{n'}$	$\bar{x}_n$	$C_x$	$C_y$	$C_{xy}$
5000	100	1.000	0.10	80	22	20	.924	1.00	.832

$$\rho_{xy} = 0.9$$

Utilising the value of C.V. for auxiliary variable, we obtain the following optimum values of  $n'$  and  $n$  and the optimum variance

$$n'=409, n=59, \text{opt. Var. } (\hat{Y}_{MRD})=38.3$$

Whereas, without using the C.V. the optimum values of  $n'$  and  $n$  and optimum variance given by Jessen [1] are  $n'=408, n=59, \text{opt Var. } (\hat{Y}_{RD})=42.7$ . Hence, the efficiency of the purposed estimator relative to the usual estimator is 111 per cent.

#### REFERENCES

- [1] Jessen R.J (1978) : *Statistical survey techniques* John Wiley and Sons, New York.
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- [3] Sukhatme P.V. and Sukhatme B.V. (1970) : *Sampling theory of surveys with application* : Indian Society of Agricultural Statistics, New Delhi.